

# Derivation of the Normal Distribution: Lecture VIII

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## Reading Question

- Assume that  $f(x) = 1$  for  $0 \leq x \leq 1$ . What is the probability density function for the transformed variable  $y = \sqrt{x}$ ?

- 1 Derivation of the Normal Distribution Function
- 2 Expected Values

# Polar Integral Form

- The trick to this integration is changing the variables into a polar form.
  - *Polar Integration:* The notion of polar integration is basically one of a change in variables. Specifically, some integrals may be ill-posed in the traditional Cartesian plane, but easily solved in a polar space.
  - By polar space, any point  $y = f(x)$  can be written in a trigonometric form

$$\begin{aligned}r &= \sqrt{x^2 + y^2} \\ \theta &= \tan^{-1}(x/y) \\ y &= r \cos(\theta) \\ x &= r \sin(\theta)\end{aligned}\tag{1}$$

# Example Polar Map

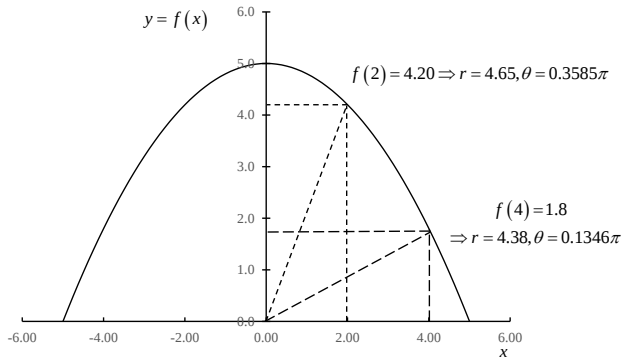
- Start with a quadratic function

$$y = f(x) = 5 - \frac{x^2}{5} \quad x \in [-5, 5] \quad (2)$$

- Consider the point  $f(4.0) = 1.8$ .
- The length of the ray from the origin is

$$r(4, 1.8) = \sqrt{4^2 + 1.8^2} = 4.38 \quad (3)$$

# Example Plot



# Inscribed Angle

- We can compute the value of the inscribed angle

$$\tan(\theta(4, 1.8)) = \frac{1.8}{4} = 0.45. \quad (4)$$

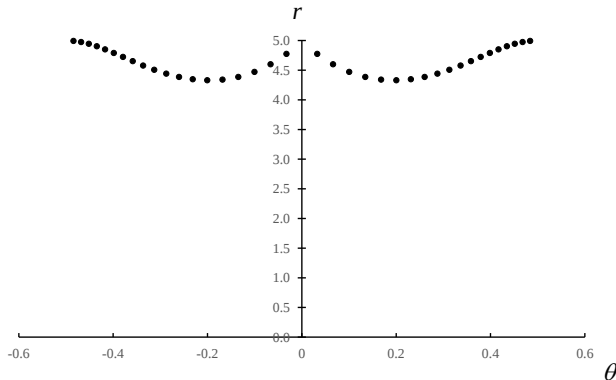
- The question is then: What angle has a  $\tan$  of 0.45

$$\theta(4, 1.8) = \tan^{-1}(0.45) = 0.1346\pi. \quad (5)$$

- Basically, we are transforming the old functional relationship  $f(x, y) = 0$  into a new relationship  $g(\theta, r(\theta))$ . Hence

$$r(\theta) = 4.38. \quad (6)$$

# Polar Transformation of Simple Quadratic



# Derivation of the Normal

- The order of proof of the normal distribution function is to start with the standard normal:

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \quad (7)$$

- First, we need to demonstrate that the distribution function does integrate to one over the entire sample space, which is  $-\infty$  to  $\infty$ . This is typically accomplished by proving the constant.

- Let us start by assuming that

$$I = \int_{-\infty}^{\infty} e^{-\frac{y^2}{2}} dy \quad (8)$$

Squaring this expression yields

$$\begin{aligned} I^2 &= \int_{-\infty}^{\infty} e^{-\frac{y^2}{2}} dy \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx \\ &= \int_{-\infty}^{\infty} e^{-\frac{y^2+x^2}{2}} dy dx \end{aligned} \quad (9)$$

- By substituting  $y = r \cos(\theta)$  and  $x = r \sin(\theta)$  we derive

$$\begin{aligned} y^2 + x^2 &= r^2 \cos^2(\theta) + r^2 \sin^2(\theta) \\ &= r^2 [\cos^2(\theta) + \sin^2(\theta)] \\ &= r^2 \end{aligned} \quad (10)$$

Defining the derivatives of the change in variables

$$dydx = r dr d\theta \quad (11)$$

Folding these two results together we get

$$I^2 = \int_0^{2\pi} \int_0^\infty r \exp\left(-\frac{r^2}{2}\right) dr d\theta = \int_0^{2\pi} d\theta = 2\pi \quad (12)$$

- Backing up a little bit - consider the derivative

$$\frac{\partial \exp\left(-\frac{r^2}{2}\right)}{\partial r} = -r \exp\left(-\frac{r^2}{2}\right). \quad (13)$$

Therefore

$$\int r \exp\left(-\frac{r^2}{2}\right) dr = -\exp\left(-\frac{r^2}{2}\right) \quad (14)$$

- Over the interval

$$\left(-\exp\left(-\frac{r^2}{2}\right)\right)\Big|_0^\infty \rightarrow -\exp\left(-\frac{\infty^2}{2}\right) + \exp\left(-\frac{0^2}{2}\right) = 1 \quad (15)$$

- Taking the square root of each side yields

$$I = \sqrt{2\pi} \quad (16)$$

Thus, we know that

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy = 1 \quad (17)$$

- The expression above is the expression for the standard normal. A more general form of the normal distribution function can be derived by defining a transformation function. Defining

$$\begin{aligned} y &= a + bx \\ x &= \frac{y - a}{b} \end{aligned} \tag{18}$$

By the change in variable technique, we have

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-a)^2}{2b^2}} \left| \frac{1}{b} \right| \tag{19}$$

# Expected Values

- **Definition 2.2.1.** The expected value or mean of a random variable  $g(X)$  denoted  $E[g(x)]$  is

$$E[g(X)] = \begin{cases} \int_{-\infty}^{\infty} g(x) f(x) dx & \text{if } x \text{ is continuous} \\ \sum_{x \in X} g(x) f_x & \text{if } x \text{ is discrete} \end{cases} \quad (20)$$

- The most general form used in this definition allows for taking the expectation of the function  $g(X)$ . Strictly speaking, the mean of the distribution is found where  $g(X) = X$ , or

$$E[x] = \int_{-\infty}^{\infty} x f(x) dx \quad (21)$$

- For example, we take the exponential distribution

$$\begin{aligned}
 E[x] &= \int_0^{\infty} \frac{1}{\lambda} x e^{-\frac{x}{\lambda}} dx \\
 &= - \left( x e^{-\frac{x}{\lambda}} \Big|_0^{\infty} + \int_0^{\infty} e^{-\frac{x}{\lambda}} dx \right) \\
 &= - \left( \lambda e^{-\frac{x}{\lambda}} \Big|_0^{\infty} \right) = \lambda
 \end{aligned} \tag{22}$$

- **Theorem 2.2.1.** Let  $X$  be a random variable and let  $a$ ,  $b$ , and  $c$  be constants. Then for any functions  $g_1(X)$  and  $g_2(X)$  whose expectations exist:
  - $E[ag_1(X) + bg_2(X)] = aE[g_1(X)] + bE[g_2(X)] + c.$
  - If  $g_1(X) \geq 0$  for all  $X$ , then  $E[g_1(X)] \geq 0.$
  - If  $g_1(X) \geq g_2(X)$  for all  $X$ , then  $E[g_1(X)] \geq E[g_2(X)].$
  - If  $a \leq g_1(X) \leq b$  for all  $X$ , then  $a \leq E[g_1(X)] \leq b.$

## Reading for Next Time

- Moss, Charles B. and J. S. Shonkwiler. 1993. Estimating Yield Distributions Using a Stochastic Trend Model and Nonnormal Errors. *American Journal of Agricultural Economics* 75(4), 1056–62. [DOI: 10.2307/1243993, JSTOR: 1243993](Google Scholar – 210, Web of Science – 69, Scopus – 84)
- Ramirez, Octavio A., Charles B. Moss, and William G. Boggess. 1994. Estimation and Use of the Inverse Hyperbolic Sine Transformation to Model Non-Normal Correlated Random Variables. *Journal of Applied Statistics* 21: 289-304. [DOI: 10.1080/757583872] (Web of Science – 10, Scopus – 15)